## Solutions to Assignment 3

1. Recall the experiment from Lecture 4 which was as follows.
2. We flip three coins.
3. If all three coins are different, we record the occurrence of events as follows: 1st event for the sequence $(T, H, H)$, 2nd event for $(H, T, H)$, 3rd event for $(H, H, T)$, 4th event for $(H, T, T)$, 5th event for $(T, H, T)$ and 6 th event for $(T, T, H)$.
4. All three coins are the same, then we go back to step 1 and continue!

Calculate the probability of getting the result $(T, H, H)$ using the limit law for probability.

Solution: Let $S_{k}$ be the event that on the $k$-flip of three coins, all are the same; we have $P\left(S_{k}\right)=2 / 8=1 / 4$. Let $T_{k}$ be the event that on the $k$-th flip of three coins we get $(T, H, H)$; we have $P\left(T_{k}\right)=1 / 8$. The event $A_{k}$ that results in $(T, H, H)$ being recorded is $S_{1} \cap \ldots S_{k-1} \cap T_{k}$. By the independence of these events, we have $P\left(A_{k}\right)=(1 / 4)^{k-1} \cdot(1 / 8)$.
The event that $(T, H, H)$ is recorded is the same as $B=\cup_{n} A_{n}$. Let $B_{n}=A_{1} \cup \ldots A_{n}$. Since $A_{n}$ are mutually exclusive events, $P\left(B_{n}\right)=\sum_{k=1}^{n}(1 / 4)^{n-1} \cdot(1 / 8)$. By the limit law

$$
P(B)=\sup _{n} P\left(B_{n}\right)=\sum_{k=0}^{\infty}(1 / 4)^{n-1} \cdot(1 / 8)=(1 / 8) \cdot \frac{1}{1-(1 / 4)}=1 / 6
$$

(Note that this is what justifies the replacement of the roll of dice by this coin experiment.)
2. We repeated roll a pair of dice until the sum is not 7 . Find the probability of each possible sum (from 2 to 12).

Solution: For the standard roll of a pair of dice, the events $S_{k}$ for the sum being $k$
correspond to pairs of rolls given below:

$$
\begin{aligned}
S_{2} & \leftrightarrow\{(1,1)\} \\
S_{3} & \leftrightarrow\{(1,2),(2,1)\} \\
S_{4} & \leftrightarrow\{(1,3),(2,2),(3,1)\} \\
S_{5} & \leftrightarrow\{(1,4),(2,3),(3,2),(4,1)\} \\
S_{6} & \leftrightarrow\{(1,5),(2,4),(3,3),(4,2),(5,1)\} \\
S_{7} & \leftrightarrow\{(1,6),(2,5),(3,4),(4,3),(5,2),(6,1)\} \\
S_{8} & \leftrightarrow\{(2,6),(3,5),(4,4),(5,3),(6,2)\} \\
S_{9} & \leftrightarrow\{(3,6),(4,5),(5,4),(6,3)\} \\
S_{10} & \leftrightarrow\{(4,6),(5,5),(6,4)\} \\
S_{11} & \leftrightarrow\{(5,6),(6,5)\} \\
S_{12} & \leftrightarrow\{(6,6)\}
\end{aligned}
$$

It follows that the probabilities are

$$
\begin{aligned}
P\left(S_{2}\right) & =1 / 36 \\
P\left(S_{3}\right) & =2 / 36 \\
P\left(S_{4}\right) & =3 / 36 \\
P\left(S_{5}\right) & =4 / 36 \\
P\left(S_{6}\right) & =5 / 36 \\
P\left(S_{7}\right) & =6 / 36 \\
P\left(S_{8}\right) & =5 / 36 \\
P\left(S_{9}\right) & =4 / 36 \\
P\left(S_{10}\right) & =3 / 36 \\
P\left(S_{11}\right) & =2 / 36 \\
P\left(S_{12}\right) & =1 / 36
\end{aligned}
$$

In the given experiment, we keep rolling until we do not get 7. Consider the event $T_{k, n}$ that we get a non- 7 for the first time on the $n$-th roll and the resulting sum is $k$. Probability of $T_{k, n}$ is the same as the probability of 7 on $n-1$ rolls followed by $k$ on the $n$-th roll. By independence of the events this is $(1 / 6)^{n-1} \cdot P\left(S_{k}\right)$. Now all the events $T_{k, n}$ are mutually exclusive. Hence the probability of the event $T_{k}=\cup_{n} T_{k, n}$ is given by

$$
P\left(T_{k}\right)=\sum_{n=1}^{\infty}(1 / 6)^{n-1} P\left(S_{k}\right)=P\left(S_{k}\right) \cdot(6 / 5)
$$

It follows that we have

$$
\begin{aligned}
P\left(T_{2}\right) & =1 / 30 \\
P\left(T_{3}\right) & =2 / 30 \\
P\left(T_{4}\right) & =3 / 30 \\
P\left(T_{5}\right) & =4 / 30 \\
P\left(T_{6}\right) & =5 / 30 \\
P\left(T_{8}\right) & =5 / 30 \\
P\left(T_{9}\right) & =4 / 30 \\
P\left(T_{10}\right) & =3 / 30 \\
P\left(T_{11}\right) & =2 / 30 \\
P\left(T_{12}\right) & =1 / 30
\end{aligned}
$$

3. Given a real-valued random variable $X$, what is the relation between $P(X \leq 0)$ and the probabilities $P(X \leq 1 / n)$ for all positive integer values of $n$. Is there anything special about 0 in this example?

Solution: If $A_{n}$ denotes the event $X \leq 1 / n$ then the intersection of all of these $A=\cap_{n} A_{n}$ is the same as $X \leq 0$. Since the $A_{n}$ are decreasing (non-increasing) it follows that

$$
P(X \leq 0)=\inf _{n} P(X \leq 1 / n)
$$

We can replace 0 by any real number $c$, and $1 / n$ by $c+1 / n$ the same reasoning works to give

$$
P(X \leq c)=\inf _{n} P(X \leq c+1 / n)
$$

4. Given a real-valued random variable $X$ what is the relation between $P(X \leq 0$ and the probabilities $P(X \leq-1 / n)$ for all positive integer values of $n$. Is there anything special about 0 in this example?

Solution: If $A_{n}$ denotes the event $X \leq-1 / n$ then the union of all of these $A=\cap_{n} A_{n}$ is the same as $X \leq 0$. Since the $A_{n}$ are increasing (non-decreasing) it follows that

$$
P(X<0)=\sup _{n} P(X \leq-1 / n)
$$

Moreover, we have $X \leq 0$ as the disjoint union of the events $X<0$ and $X=0$, hence we have

$$
P(X \leq 0)=P(X=0)+\sup _{n} P(X \leq-1 / n)
$$

We can replace 0 by any real number $c$, and $1 / n$ by $c+1 / n$ the same reasoning works to give

$$
P(X \leq c)=P(X=c)+\sup _{n} P(X<c-1 / n)
$$

5. In a repeatable experiment $\Sigma$ there are events $A$ and $B$ which can be observed. Assume that $P(B)>0$. We now carry out $\Sigma$ repeatedly until $B$ is observed. What is the probability that $A$ is also observed at the same time as $B$ ?

Solution: Let $T_{n}$ denote the event that $A \cap B$ is observed on the $n$-th repetition of $\Sigma$ while $B^{c}$ is observed on all $n-1$ earlier repetitions of $\Sigma$. The probability of $T_{n}$ is $P\left(B^{c}\right)^{n-1} P(A \cap B)=(1-P(B))^{n-1} P(A \cap B)$. Let $T$ be the event that $A \cap B$ is observed after repeatedly observing $B^{c}$; this is the union of $T_{n}$ over all $n$. Since the events $T_{n}$ are mutually exclusive, we see that

$$
P(T)=\sum_{n=1}^{\infty}(1-P(B))^{n-1} P(A \cap B)=P(A \cap B) / P(B)=P(A \mid B)
$$

Thus, this repeated experimentation until we see $B$ is an interpretation of conditional probability of $A$ given $B$.
6. We know that $40 \%$ of the fish in a pond are female. We repeatedly catch and release fish until we find one that is female; in that case we record its species before releasing it. After 100 such recordings we find that 20 of these females were pomfret fish. What is a reasonable estimate of the percentage of female pomfret in the pond? Can we conclude the $20 \%$ of the fish are pomfret?

Solution: Let $P(F P)$ denote the probability of a randomly chosen fish being a female pomfret. The probability that we pick such a fish on the $n$-th trial is the same as the probability that we pick $n-1$ male fish followed by this fish; in other words it is $(1-0.4)^{n-1} P(F P)$. It follows that the probability of recording a pomfret by the mechanism used is $\sum_{n}(0.6)^{n-1} P(F \cap P)=(1 / 0.4) P(F P)$. By the above frequency, of $20 / 100$ we can estimate $0.2=(1 / 0.4) P(F P)$ or $P(F P)=0.08$. In other words, we estimate that $8 \%$ of the fish are female pomfret.
We cannot conclude that $20 \%$ of the fish are pomfret. We must also do a similar experiment where we only record the species of male fish, suppose that in this case
$30 \%$ of the fish turn out to be pomfret. From a calculation like the one above we will get the probability of male pomfret as 0.18 . Putting these together we would get the percentage of pomfret as $26 \%$.

